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1 Proof of Claim 2

Claim. Given a fixed setting of the parameters θ there exists some $0 \leq \lambda < 1$ such that for each M-statistic the difference between the true value and the value as approximated by the ASOS procedure can be expressed as a linear function of the approximation error in the DAS whose operator norm is bounded above by $ck_{lim}^2 \lambda^{k_{lim}-1}$ for some constant c that doesn't depend on k_{lim} .

Proof. (sketch) We will use the Δ symbol to denote the error in a given term. So for example, $\Delta(y, x^*)_{k_{lim}+1}$ will denote error due to in the first DAS.

By repeated application of the first 2nd-order recursion we have that $\Delta(y, x^*)_k = \Delta(y, x^*)_{k_{lim}+1}(H')^{k_{lim}+1-k}$. Then by repeated application of the second 2nd-order recursion we have that $\Delta(x^*, y)_k = H^k \Delta(y, x^*)'_0 = H^{k+k_{lim}+1}\Delta(y, x^*)_{k_{lim}+1}'$. We can already see a pattern starting to emerge here. The error for statistics of small k-value/time-lag is given by the approximation error in the DAS, multiplied by some large power of H. If H is "small" in some sense then the error will decay exponentially as k decreases. This will be made more rigorous later.

Repeatedly applying the third and fifth 2nd-order recursions we also have:

$$\Delta(x^*, x^*)_k = \Delta(x^*, x^*)_{k_{lim}} (H')^{k_{lim}-k} + \sum_{i=k}^{k_{lim}} \Delta(x^*, y)_i K'(H')^{i-k}$$
$$\Delta(x^T, y)_k = J^{k_{lim}-k} \Delta(x^T, y)_{k_{lim}} + \sum_{i=k}^{k_{lim}} J^{i-k} P \Delta(x^*, y)_i$$

If we plug in the previously derived error formulae into these (for the terms $\Delta(x^*, y)_i$) we note that at each term being summed is multiplied by H and/or J a total of $k_{lim} - k$ times or more. Repeated application of the sixth 2nd-order equation gives:

$$\Delta(x^T, x^*)_k = J^{k_{lim}-k} \Delta(x^T, x^*)_{k_{lim}} + \sum_{i=k}^{k_{lim}} J^{i-k} P \Delta(x^*, x^*)_i$$

The right hand side of this equation contains terms of the form $\Delta(x^*, x^*)_i$, multiplied by J^{i-k} . Thus as before, the combined power H and J in each term of the sum is $\geq k_{lim} - i + i - k = k_{lim} - k$, and there are roughly k_{lim}^2 such terms.

The spectrums of J and H are equal (basic LDS result) and their spectral radius λ (the maximum of the magnitudes of the eigenvalues) is less than 1. In practice we have found that λ is often significantly less than 1 even when the spectral radius of A is relatively close to 1. Intuitively J and H capture the strength of the dependency between the hidden states in consecutive time-steps. Smaller eigenvalues correspond to eigen-components with weaker dependencies that decay faster. Letting $\sigma(X)$ denote the spectral radius of an arbitrary matrix X and using the basic property that $\sigma(XY) \leq \sigma(X)\sigma(Y)$ and the identity $\|B\| \leq \dim(B)\sigma(B)$ we can estimate the 2-norms of various error matrices in terms of the 2-norms of the DAS errors. For example, we have that $\|\operatorname{vec}(\Delta(y, x^*)_k))\|_2 \leq N_x^2 \|\Delta(y, x^*)_{k_{lim}+1}\|_2 \lambda^{k_{lim}+1-k}$. For harder cases such as $\Delta(x^T, y)_k$ that involve the sum over many terms

For harder cases such as $\Delta(x^T, y)_k$ that involve the sum over many terms we can apply triangle inequality for norms and then bound the norm of each term. This is the reason that the factor k_{lim}^2 appears in the claimed operator norm bound.

The most difficult case is $\Delta(x^T, x^T)_0$. The ASOS procedure estimates $(x^T, x^T)_0$ by solving the 2nd ASOS equation as a Lyapanov equation. The resultant error in $(x^T, x^T)_0$ is thus also the solution of a similar Lyapanov equation:

$$\Delta(x^{T}, x^{T})_{0} = J\Delta(x^{T}, x^{T})_{0}J' + J\Delta(x^{T}, x^{*})_{1}P' + P\Delta(x^{T}, x^{*})_{0}'$$

While it is possible, although unlikely, that solving this equation could greatly amplify the error, this effect would be linear and constant (since the linear coefficients on $(x^T, x^T)_0$ do not depend on k_{lim}) and thus bounded in norm.

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2 Proof of Claim 3

Claim. For i = 1, 2, 3:

$$\lim_{T \to \infty} \mathcal{E}_{\theta} \left[\left\| \frac{1}{T} \operatorname{vec}(\phi_i) \right\|_2^2 \right] = 0$$

Proof. First we will consider the case i = 1.

Define residual prediction error γ_t by $\gamma_t = y_t - E_{\theta} [y_t | y_{\leq t-1}]$ and note that $\phi_1 \equiv (y, x^*)_{k_{lim}+1} - CA ((x^*, x^*)_{k_{lim}} - x_T^* x_{T-k_{lim}}^*)$ can be expressed as $\sum_{t=1}^{T-k-1} \gamma_{t+k+1} x_t^{*'}$.

Using this fact and the linearity of expectation we have that the expectation can be written as:

$$\begin{aligned} \mathbf{E}_{\theta} \left[\| \frac{1}{T} \operatorname{vec}(\phi_{1}) \|_{2}^{2} \right] &= \mathbf{E}_{\theta} \left[\frac{1}{T^{2}} \operatorname{tr}(\phi_{1} \ \phi_{1}') \right] \\ &= \frac{1}{T^{2}} \operatorname{tr}(\sum_{t=1}^{T-k-1} \sum_{s=1}^{T-k-1} \mathbf{E}_{\theta} \left[\operatorname{vec}(\gamma_{t+k+1} x_{t}^{*\prime}) \operatorname{vec}(\gamma_{s+k+1} x_{s}^{*\prime})' \right]) \\ &= \frac{1}{T^{2}} \operatorname{tr}(\sum_{t=1}^{T-k-1} \sum_{s=1}^{T-k-1} \mathbf{E}_{\theta} \left[x_{t}^{*} x_{s}^{*\prime} \otimes \gamma_{t+k+1} \gamma_{s+k+1}' \right]) \end{aligned}$$

First we consider the terms of the inner sum where $t \neq s$. By symmetry we may assume, without loss of generality, that s > t. Then using the law of iterated expectation and the fact that $\forall i \ E_{\theta}[\gamma_i] = 0$ we have:

$$\begin{aligned} \mathbf{E}_{\theta} \left[x_t^* x_s^{*'} \otimes \gamma_{t+k+1} \gamma_{s+k+1}' \right] &= \mathbf{E}_{\theta} \left[\mathbf{E}_{\theta} \left[x_t^* x_s^{*'} \otimes \gamma_{t+k+1} \gamma_{s+k+1}' \mid y_{\leq s+k} \right] \right] \\ &= \mathbf{E}_{\theta} \left[x_t^* x_s^{*'} \otimes \gamma_{t+k+1} \otimes \mathbf{E}_{\theta} \left[\gamma_{s+k+1}' \mid y_{\leq s+k} \right] \right] \\ &= \mathbf{E}_{\theta} \left[x_t^* x_s^{*'} \otimes \gamma_{t+k+1} \otimes \mathbf{0} \right] = \mathbf{0} \end{aligned}$$

For terms where t = s we have instead that:

where we recall that $S_i \equiv \operatorname{Cov}_{\theta}[\gamma_t \mid y_{\leq t+k}] = \operatorname{E}_{\theta}[\gamma_t \gamma'_t \mid y_{\leq t+k}]$ Our final equation for the expectation is then:

$$\frac{1}{T^2} \operatorname{tr}(\sum_{t=1}^{T-k-1} \operatorname{E}_{\theta}[x_t^* x_t^{*'}] \otimes S_{t+k+1}) = \frac{1}{T^2} \sum_{t=1}^{T-k-1} \operatorname{E}_{\theta}[\|x_t^*\|_2^2] \cdot \operatorname{tr}(S_{t+k+1})$$

Intuitively, the growth of the sum is linear in T, not quadratic, and thus the factor $\frac{1}{T^2}$ will cause the entire right-hand expression to go to zero in the limit. We can formalize this intuition. This first tool we will need as a basic result about the asymptotic behavior of the LDS: under the control-theoretic conditions necessary for steady-state we also have that distributions over x_t^* and y_t approach equilibrium as $t \to \infty$.

A simple consequence of this result is that various expectations over x_t^* and y_t will converge as $t \to \infty$. Thus we have $\lim_{T\to\infty} \frac{1}{T} \sum_{t=1}^{T-k-1} \mathbb{E}_{\theta}[x_t^* x_t^{*'}] = X_0$ where $X_0 = \lim_{t\to\infty} \mathbb{E}_{\theta}[x_t^* x_t^{*'}]$. We also know that S_t approaches its steady-state value S as as $t \to \infty$. These two facts allow us to evaluate the limit:

$$\lim_{T \to \infty} \frac{1}{T^2} \sum_{t=1}^{T-k-1} \mathcal{E}_{\theta} \left[\|x_t^*\|_2^2 \right] \cdot \operatorname{tr}(S_{t+k+1}) = \lim_{T \to \infty} \frac{1}{T} \operatorname{tr}(X_0) \operatorname{tr}(S) = 0$$

For the remain cases of i we can show, using a similar argument to the one given above, that:

$$E_{\theta}\left[\|\frac{1}{T}\operatorname{vec}(\phi_{2})\|_{2}^{2}\right] = \frac{1}{T^{2}}\sum_{t=1}^{T-k-1} E_{\theta}\left[\|x_{t}^{*}\|_{2}^{2}\right] \cdot \operatorname{tr}(V_{t+k,t+k}^{t+k-1})$$
$$E_{\theta}\left[\|\frac{1}{T}\operatorname{vec}(\phi_{3})\|_{2}^{2}\right] = \frac{1}{T^{2}}\sum_{t=1}^{T-k-1} E_{\theta}\left[\|y_{t}\|_{2}^{2}\right] \cdot \operatorname{tr}(V_{t+k,t+k}^{t+k-1})$$

and that these expectations also converge to 0 in the limit as $T \to \infty$.

3 Proof of claim 4

Claim. The approximation error in $\frac{1}{T}$ -scaled 2nd-order statistics as estimated by the ASOS procedure converges to 0 in expected squared $\|\cdot\|_2$ -norm as $T \to \infty$.

Lemma 1. Let X be a large vector formed by concatenating the vectorizations of the true values of all the 2nd-order statistics that are estimated at some point during the ASOS procedure (this includes the DAS, the M-statistics, and all of the intermediate quantities). Let \hat{X} be the corresponding approximate estimate obtained from the ASOS procedure. Then we have:

$$\lim_{T \to \infty} \mathbf{E}_{\theta} \left[\| \frac{1}{T} (X - \hat{X}) \|_{2}^{2} \right] = 0$$

Proof. The system of equations solved by the ASOS procedure consists of the ASOS equations, the ASOS approximations, and of a particular *subset* of the 2nd-order equations. All of these are linear in the 2nd-order statistics. Moreover, except for the ASOS approximations they are all satisfied by the exact values of the second-order statistics. We may thus write the system as:

$$\Upsilon X = \Gamma + \Phi$$

where Υ is a matrix of coefficients, Γ a vector that accounts for the constant terms in each equation (i.e. those only involving statistics of the form $(y, y)_k$ for some k) and Φ is a vector that accounts for errors in the ASOS approximations.

The ASOS procedure is simply a computationally efficient method for solving this system where the unknown Φ is replaced by the zero vector. So, $\hat{X} = \Upsilon^{-1}\Gamma$ while the true value is given by $X = \Upsilon^{-1}(\Gamma + \Phi)$. Thus $\frac{1}{T}(X - \hat{X}) = \Upsilon^{-1}(\frac{1}{T}\Phi)$ and hence the expectation in the claim can be rewritten and then bounded:

$$\mathbf{E}_{\theta} \left[\| \Upsilon^{-1}(\frac{1}{T}\Phi) \|_{2}^{2} \right] \leq \| \Upsilon^{-1} \|^{2} \mathbf{E}_{\theta} \left[\| \frac{1}{T}\Phi \|_{2}^{2} \right]$$

But by the previous claim $\lim_{T\to\infty} E_{\theta} \left[\|\frac{1}{T}\Phi\|_2^2 \right] = 0$ and then noting that Υ doesn't depend on T, the result follows.

4 Solving The "Primary Equation"

Lemma 1. Let V be a vector space, $f : V \longrightarrow V$ be a continuous linear function such that $\rho(f) < 1$. Then a solution to the equation x = f(x) + y is given by:

$$x_0 = \sum_{i=0}^{\infty} f^i(y) \tag{1}$$

where the exponents denote function composition.

Proof. The condition $\rho(f) < 1$ ensures that the series converges (and determines the rate of convergence).

Then,

$$x_0 = \sum_{i=0}^{\infty} f^i(y) = \sum_{i=1}^{\infty} f^i(y) + f^0(y) = \sum_{i=0}^{\infty} f \circ f^i(y) + y = f(\sum_{i=0}^{\infty} f^i(y)) + y = f(x_0) + y$$

Algorithm 1 Algorithm solving the primary equation

1: Input: A, C, K, H, G 2: Initialize X := 0, Y := G. 3: while Y has not converged to 0 do 4: Y := Solution for Z of (Z = AZH' + Y)5: X := X + Y6: $Y := H^{2k_{lim}+1}Y'A'C'K'$ 7: end while

where we have used the fact that f is both continuous and linear so that it respect the infinite sum.

Now let $f_1(X) = X - AXH'$, $f_2(X) = H^{2k_{lim}+1}X'A'C'K'$ and y = Gwhere A, H, C, K and G. These functions are clearly linear in X and continuous. Then the solution of $f_1(X) = f_2(X) + G$ is the solution of the primary equation. Taking f_1^{-1} of both sides yields $X = f_1^{-1} \circ f_2(X) + f_1^{-1}(G)$ which is the form of the equation solved in the previous lemma with $f = f_1^{-1} \circ f_2$ and $y = f_1^{-1}(G)$.

Conjecture 1. For all $k_{lim} \ge 0$, $f_1^{-1} \circ f_2$ is a continuous linear function with $\rho(f_1^{-1} \circ f_2) < 1$

In practice, $\rho(f_1^{-1} \circ f_2)$ will be a significantly less than 1 when k_{lim} is large enough (even when $\rho(H)$ is close to 1) which implies rapid convergence of the series defined in (1).

Algorithm 1 computes this series term-by-term and so by the previous lemmas and rapid converge property it is an efficient method for solving the primary equation. Note that Y can be computed easily on line 4 because Z = AZH' + Y is a Sylvester equation, for which there are know efficient algorithms.

5 Psuedo-code for ASOS

Algorithm 2 The ASOS algorithm for computing the E-step. Note that for the purposes of implementation, symbols such as $(y, x)_k^{\dagger}$ can simply be interpreted as k-indexed matrix-valued variable names.

1: perform steady-state computations (Algorithm 3) 2: compute approximate first and last k_{lag} 1st-order statistics (Algorithm 4) 3: $(y, x^*)_{k_{lim}+1}^{\dagger} := -CAx_T^* x_{T-k_{lim}}^*$ 4: **for** $k = k_{lim}$ down to 0 **do** $(y, x^*)_k^{\dagger} := (y, x^*)_{k+1}^{\dagger} H' + ((y, y)_k - y_{1+k} y_1') K' + y_{1+k} x_1^{*'}$ 5:6: end for 7: $(x^*, y)_0^{\dagger} := (y, x^*)_0^{\dagger'}$ 8: for k = 1 to k_{lim} do $(x^*, y)_k^{\dagger} := H((x^*, y)_{k-1}^{\dagger} - x_T^* y_{T-k+1}') + K(y, y)_k$ 9: 10: end for 11: $G := (-Ax_T^* x_{T-k_{lim}}^* + ((x^*, y)_{k_{lim}}^\dagger - x_{1+k_{lim}}^* y_1')K' + x_{1+k_{lim}}^* x_1^{*'}$ 12: $(x^*, x^*)_{k_{lim}} :=$ SolvePrimaryEquation(A, C, K, H, G)13: for $k \in \{0, 1, 2, ..., k_{lim}\}$ do $(y, x^*)_k := (y, x^*)_k^{\dagger} + CA(x^*, x^*)_{k_{lim}} H^{k_{lim}+1-k'}$ 14: $(x^*, y)_k := (x^*, y)_k^{\dagger} + H^{k_{lim}+1+k}(x^*, x^*)'_{k_{lim}}A'C'$ 15:16: end for 17: for $k = k_{lim}$ down to 0 do $(x^*, x^*)_k := (x^*, x^*)_{k+1} H' + \left((x^*, y)_k - x^*_{1+k} y'_1 \right) K' + x^*_{1+k} x^*_1$ 18:19: **end for** 20: $(x^T, x^*)_{k_{lim}} := (x^*, x^*)_{k_{lim}}$ 21: for $k = k_{lim} - 1$ down to 0 do $(x^{T}, x^{*})_{k} := J(x^{T}, x^{*})_{k+1} + P\left((x^{*}, x^{*})_{k} - x_{T}^{*} x_{T-k}^{*}\right) + x_{T}^{T} x_{T-k}^{*}$ 22: 23: end for 24: $(x^*, x^T)_0 := (x^T, x^*)_0'$ 25: $L := -Jx_1^T x_1^{T'} J' + J(x^T, x^*)_1 P' + P((x^*, x^T)_0 - x_T^* x_T^{T'}) + x_T^T x_T^{T'}$ 26: $(x^T, x^T)_0 :=$ SolveLyapunov(J, J', L)27: $(x^T, x^T)_1 := ((x^T, x^T)_0 - x_1^T x_1^T) J' + (x^T, x^*)_1 P'$ 28: $(x^T, y)_{k_{lim}} := (x^*, y)_{k_{lim}}$ 29: for $k = k_{lim} - 1$ down to 0 do $(x^{T}, y)_{k} := J(x^{T}, y)_{k+1} + P\left((x^{*}, y)_{k} - x_{T}^{*}y_{T-k}'\right) + x_{T}^{T}y_{T-k}'$ 30: 31: end for 32: $(y, x^T)_0 := (x^T, y)'_0$

Algorithm 3 Steady-state Computations

1: $\Lambda_0^1 := \text{SolveDARE}(A, C, Q, R)$ 2: $S := C\Lambda_0^1 C' + R$ 3: $K := \Lambda_0^1 C' S^{-1}$ 4: $\Lambda_0^0 := \Lambda_0^1 - KC\Lambda_0^1$ 5: $J := \Lambda_0^0 A' (\Lambda_0^1)^{-1}$ 6: $\Lambda_0 := \text{SolveSylvester}(J, J', \Lambda_0^0 - J\Lambda_0^1 J')$ 7: $V_1^T := V_0^T J'$ 8: H := A - KCA9: P := I - JA

Algorithm 4 Compute Approximate First and Last k_{lag} 1st-order Statistics

1: $x_1^* := \pi_1 + K(y_1 - C\pi_1)$ 2: for k = 2 to k_{lag} do 3: $x_t^* := Hx_{t-1}^* + Ky_t$ 4: end for 5: $x_{k_{lag}}^T := x_{k_{lag}}^*$ 6: for $k = k_{lag} - 1$ down to 1 do 7: $x_t^T := Jx_{t+1}^T + Px_t^*$ 8: end for 9: $x_{T-k_{lim}}^* := 0$ 10: for $k = T - k_{lag} + 1$ to T do 11: $x_t^* := Hx_{t-1}^* + Ky_t$ 12: end for